# "Pushing on a String: U.S. Monetary Policy Is Less Powerful in Recessions" by Silvana Tenreyro and Gregory Thwaites (LSE, BoE)

**Discussion by** 

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# Overview

- Paper presents evidence that the effect of monetary policy shocks on output, expenditure volumes, inflation, etc., is **regime dependent**: it is more powerful in expansions than in recessions (U.S. data)
- Monetary policy shocks to the federal funds rate are taken from Romer and Romer (2004): "narrative" approach to identification.
- Impulse responses estimated by an extension of the local projection method in Jorda (2005), extended by Auerbach and Gorodnichenko (2011):
  - no need to specify a full-fledged dynamic model (VAR)
  - extension consists of allowing for different impulse responses across two regimes (booms vs. recessions); smooth transition btw regimes
- Alternative explanation based on differential effects of positive and negative shocks examined (and ruled out)
- Array of robustness checks

#### Is the evidence really "statistically strong"?

- Let  $\beta_h^b$  and  $\beta_h^r$ , h = 0, 1, 2, ..., 20 denote impulse responses in booms vs. recessions
- We are interested in testing  $H_0^h: \beta_h^b \beta_h^r = 0, h = 0, 1, 2, \dots, 20$
- The authors consider a sequence of t-stats:  $\hat{t}_h$ ,  $h=0,1,2,\ldots,20$
- Each t-stat is compared with the (asymptotic) 5% critical values  $\pm 1.65$
- BUT: this is a **joint** testing problem—individual significance at a given level is misleading

E.g. if the t-stats were mutually independent, then

$$\begin{split} & \mathsf{Prob.}\left[ \left| \hat{t}_h \right| > 1.65 \text{ for some } h \left| H_0^0, \dots, H_0^{20} \text{ all true} \right] \\ &= 1 - 0.9^{21} = 0.89 \end{split}$$

#### Is the evidence really "statistically strong"? (cont'd)

- In this problem the t-stats are correlated—it is not immediately clear how distorted nominal significance levels are
- Still, the fact that the sequence  $\hat{t}_h$ , h = 0, 1, 2, ..., 20 crosses  $\pm 1.65$  once or twice is not the most convincing evidence against  $H_0$ .
- Jorda (2005) does discuss joint inference for  $\beta_h$ , h = 0, 1, 2, ..., H. Can you generalize to differences across regimes?
- Alternatively, you could try using the White (2000) "reality check" or the Hansen (2005) refinement.

 $\Rightarrow$  Inference based on  $\max_h \hat{t}_h$  via block-bootstrap

• (Or don't do anything but acknowledge the issue.)

## Interpretation of state dependence

- Interpretation based on asymmetric effect of positive vs. negative shocks ruled out:
  - positive (contractionary) policy shocks found to have larger effects than negative (expansionary) shocks
  - but positive shocks do not appear more common in boom periods (as defined in the paper)
- This type of asymmetry is in line with recent findings by Angrist, Kuersteiner and Jorda (2013); may want to point this out. Second part is a bit surprising.
- Policy implications of results. Is monetary stimulation by cutting the interest rate during recessions doubly doomed?
  - A: trying to do something that doesn't work well in general at a time when it is especially ineffective; or
  - B: only means that negative **surprises** are ineffective during recessions; the anticipated component of a rate cut may still be stimulative

## **Specification and estimation issues**

- Parameters that control the definition of boom vs. bust and the speed of transition are calibrated rather than estimated
  - state variable=centered seven quarter moving average of real GDP growth; 20% of sample designated as recession
- In fairness, these parameters hard to estimate (identification is weak) and the authors perform some robustness checks; is it enough?
- There are examples of Threshold-VAR models where the threshold is estimated by (quasi-)ML:
  - $\Rightarrow$  E.g., Deak and Lenarcic (2013): fiscal multiplier depends on debt-to-GDP ratio
  - $\Rightarrow$  Specification search over empirically most relevant state variable
- (This would however reintroduce the disadvantages of the VAR methodology relative to local projections)

#### **Miscellaneous comments**

- Briefly mention possible theoretical explanations of why the effect of monetary policy shocks can be expected to depend on the state of the business cycle
- Clarify role of "control" variables  $x_t$  in the projection method. Do we need them if the Romer and Romer shocks are truly exogenous?
  - In Jorda (2005) there is a vector  $y_t$  of dependent variables and  $x_t = (y_{t-1}, \dots, y_{t-p})$
  - Even if we are interested in the IRF of a single component of  $y_t$  only, what other variables to include is still a choice
  - $\boldsymbol{p}$  is still a choice
  - Robustness checks w.r.t. the specification of  $x_t$  to see if it matters

### Miscellaneous comments (cont'd)

- Extending robustness checks: use SVAR-based monetary policy shocks instead of Romer and Romer.
  - This is already done to a limited extent, but in a very minimalistic VAR
  - You could use one of the benchmark identification schemes in Christiano, Eichenbaum and Evans (1999)
  - (somewhat weird mixture between VAR modeling and local projections)