

Incomplete pass-through in import markets and permanent vs. transitory exchange-rate shocks

2nd Workshop on Macroeconomic Policy
Research, Magyar Nemzeti Bank,
October 2-3, 2003

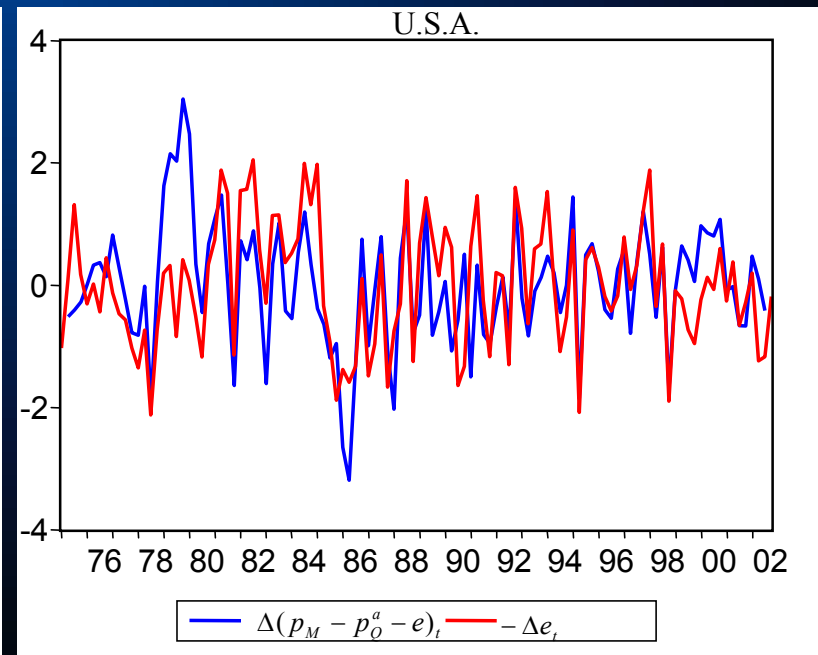
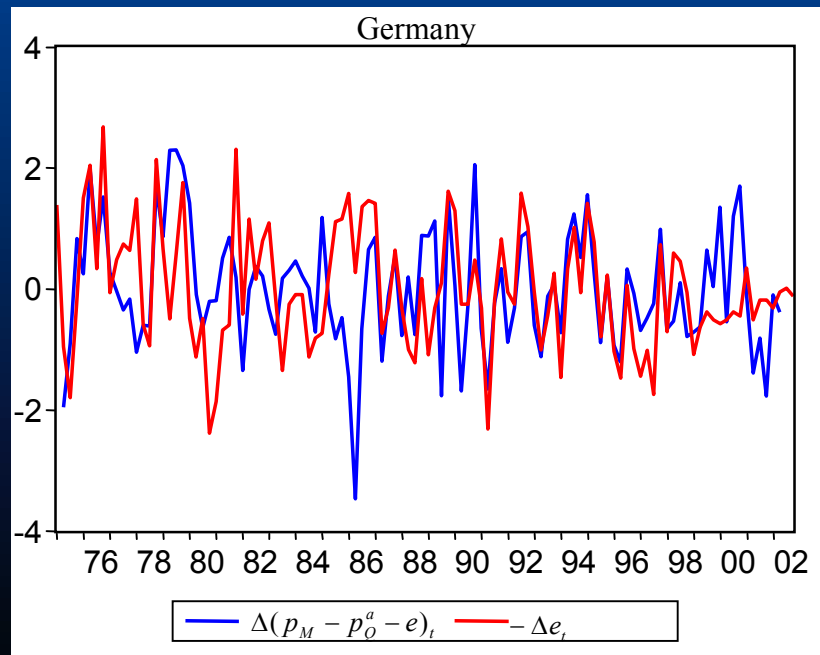
Martin Meurers, *ifo Institute for Economic Research, Munich,
Germany*

- First-glance empirical evidence for incomplete pass-through (pricing-to-market) and economic rationale
- Estimation of import supply functions for the U.S., Japan, Germany, France, and Italy
- Permanent and transitory exchange-rate changes in the context of dynamic optimizing behavior
- Decomposition of real exchange rates into a permanent and a transitory component
- Re-estimation of import supply functions with decomposed exchange rates
- Conclusion

Changes in the value of a country's currency do not lead to congruent changes of import prices

Latest empirical evidence: e.g. Campa and Goldberg (2002), Pollard and Coughlin (2003)

⇔ Import prices in relation to foreign costs in common currency [$P_M/(P_Q^a \cdot E)$] vary with the exchange rate (E).



Static PTM: Dornbusch (1987), Krugmann (1987)

$$P_M^f = \frac{\varepsilon}{1 + \varepsilon} C^{f'}(M, P_Q^a) \quad \varepsilon = \varepsilon(P_Q / \bar{P}_M^f \cdot E)$$

$$(1) \quad P_M^f = \left[\frac{P_Q}{P_M^f \cdot E} \right]^\psi \cdot C^{f'}(M, P_Q^a) \quad r = e + p_Q^a - p_Q$$

$$(2) \quad p_M - p_Q = \gamma_0 + (1 - \gamma_1)r$$

Dynamic PTM: Baldwin (1988), Froot and Klemperer (1989), Kasa (1992)

$$(3) \quad L_t = E_t \sum_{i=0}^{\infty} \theta^i [\lambda (p_{t+i} - \tilde{p}_{t+i})^2 + \Delta p_{t+i}^2] \quad p_t = p_{M_t} - p_{Q_t}$$

$$E_t \tilde{p}_{t+i} = (1 - \gamma_1) E_t r_{t+i}$$

$$(7) \quad \Delta p_t = -(1 - \mu)[p_{t-1} - (1 - \gamma_1)r_{t-1}] + t' G(\theta, \mu, \Pi, A) \Delta z_t$$

Oil price as additional determinant of import prices:

Table 1: Pair-wise correlations of oil prices with exchange rates and import prices

	<i>U.S.</i>	<i>Japan</i>	<i>Germany</i>	<i>France</i>	<i>Italy</i>
$Corr(\Delta r_t, \Delta p_{Oil_t})$	0.00	0.10	0.04	0.17	0.08
$Corr(\Delta p_t, \Delta p_{Oil_t})$	0.54	0.43	0.44	0.38	0.43

Unit-Root tests:

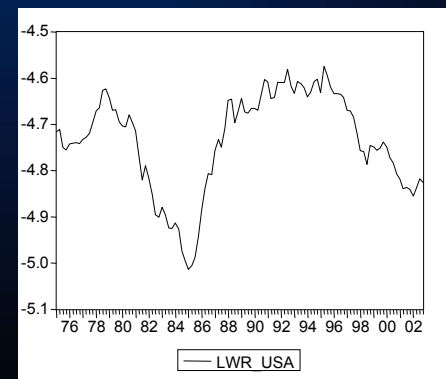
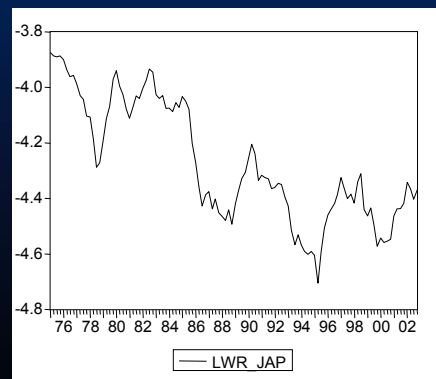
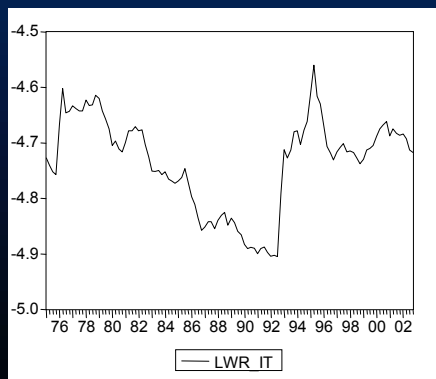
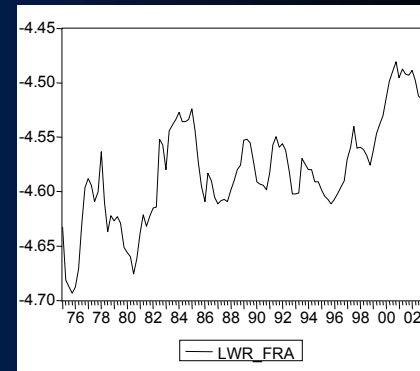
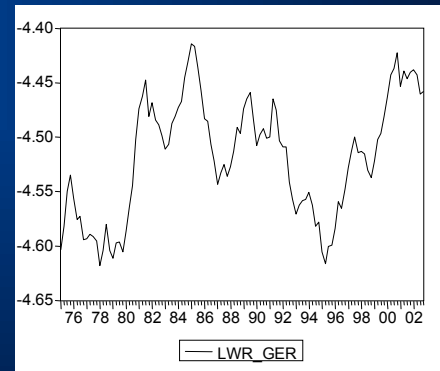
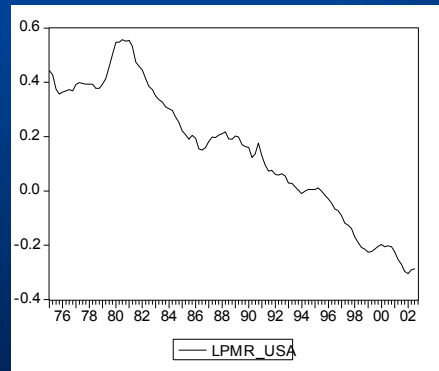
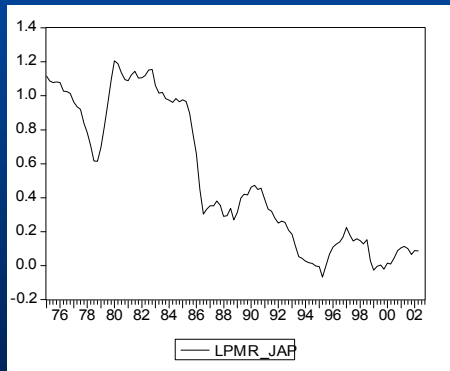
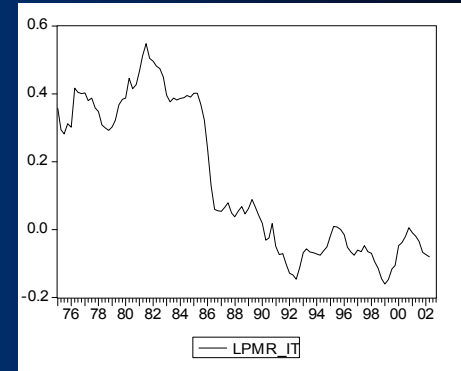
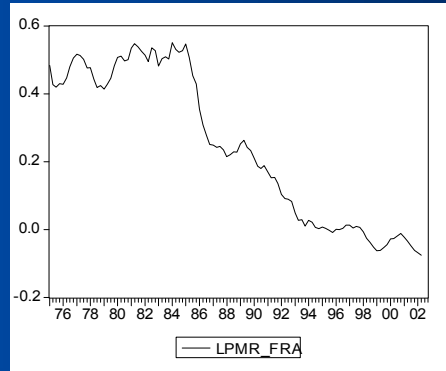
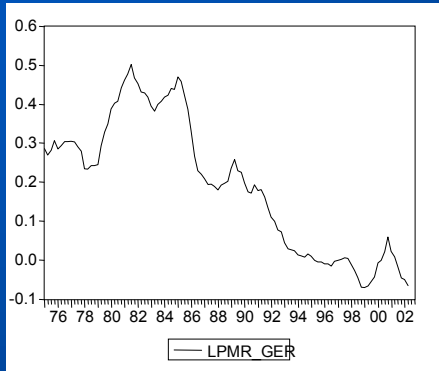
Augmented Dickey-Fuller: $H_0: \alpha = 1$

$$\Delta y_t = c_0 + (\alpha - 1)y_{t-1} + c_1 t + \sum_{i=1}^{k-1} \beta_i \Delta y_{t-i} + \varepsilon_t \quad \varepsilon_t \sim IID(0, \sigma^2)$$

Kwiatkowski et al.: $H_0: d = 0$

$$y_t = \alpha + \beta t + d \sum_{i=1}^t u_i + \varepsilon_t \quad LM = \sum_{t=1}^T S_t^2 / \hat{\sigma}_\varepsilon^2 \quad S_t = \sum_{i=1}^t e_i$$

Relative import prices and real exchange rates



Results of the unit root tests

Table 1: Unit root tests for the real exchange rates

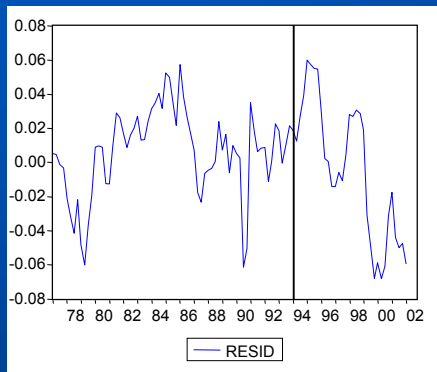
variable	Lags (k)	det. comp. in test	ADF	ADF 1 st diff.	KPSS ($l = 8$)	KPSS 1 st diff. ($l = 8$)
R_GER	4	C,T	-3.30*	-5.03***	0.128*	0.083
R_FRA	5	C,T	-3.31*	-5.32***	0.074	0.074
R_ITA	4	C,T	-2.46	-5.55***	0.157**	0.123
R_USA	5	C,T	-2.61	-3.55***	0.105	0.050
R_JAP	6	C,T	-2.70	-4.98***	0.133*	0.199

Table 2: Unit root tests for the relative import prices and the oil price

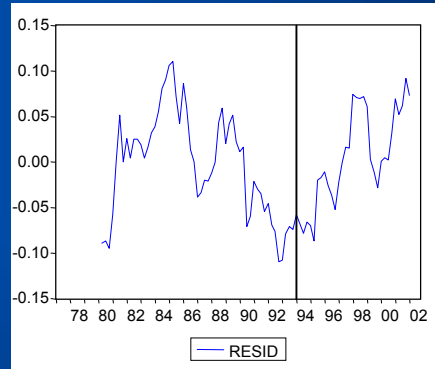
variable	Lags (k)	det. comp. in test	ADF	ADF 1 st diff.	KPSS ($l = 8$)	KPSS 1 st diff. ($l = 8$)
P_GER	6	C,T	-2.25	-3.74***	0.180**	0.147
P_FRA	5	C,T	-2.58	-4.15***	0.157**	0.118
P_ITA	7	C,T	-1.81	-3.96***	0.147**	0.096
P_USA	7	C,T	-3.25*	-3.85***	0.136*	0.142
P_JAP	7	C,T	-2.53	-4.14***	0.111	0.072
POIL	1	C	-1.37	-5.23***	0.117	0.109

Structural shifts in the European supply functions

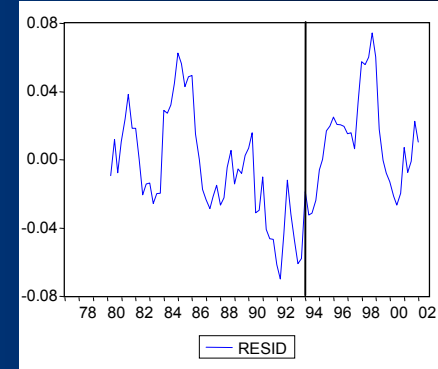
Germany



Italy



France



Summary of ECM estimations: sample 1975:1-2002:2, $T=110$

Significance of $(1-\mu)$ suggests cointegrating import supply relations – except for France

	long run ($1-\gamma_1$)	short run
U.S.	0.48	0.28
Japan	0.86	0.75
Germany	0.95	0.65
France	0.92	0.69
Italy	0.83	0.50

- Coherence with results of previous studies
- Hypothesis: different response to permanent and transitory exchange-rate changes matters

Summary of ECM estimations: sample 1975:1-2002:2, T=110

Significance of $(1-\mu)$ suggests cointegrating import supply relations – except for France

	long run ($1-\gamma_1$)	short run
U.S.	0.48	0.28
Japan	0.86	0.75
Germany	0.95	0.65
France	0.92	0.69
Italy	0.83	0.50

- Coherence with results of previous studies
- Hypothesis: different response to permanent and transitory exchange-rate changes matters

A simple model for the decomposition of real exchange rate into a permanent component q_t and a transitory component s_t

$$(10a) \quad r_t = q_t + s_t$$

$$(10b) \quad q_t = q_{t-1} + u_t \quad (s_t, u_t) \sim \text{iid}[0, \text{diag}(\sigma_s^2, \sigma_u^2)]$$

$$\rightarrow \Delta r_t = u_t + \Delta s_t$$

Prediction of the target path in the adjustment cost model:

$$E_t r_{t+i} = r_t - s_t$$

Corresponding ECM for dynamic import supply function:

$$(11) \quad \Delta p_t = -(1 - \mu)[p_{t-1} - (1 - \gamma_1)r_{t-1}] + (1 - \mu)(1 - \gamma_1)\Delta r_t - (1 - \mu)\theta\mu(1 - \gamma_1)s_t$$

Implications for conventional estimates:

- a) potential bias in estimated pass-through coefficients
- b) imprecise forecasts

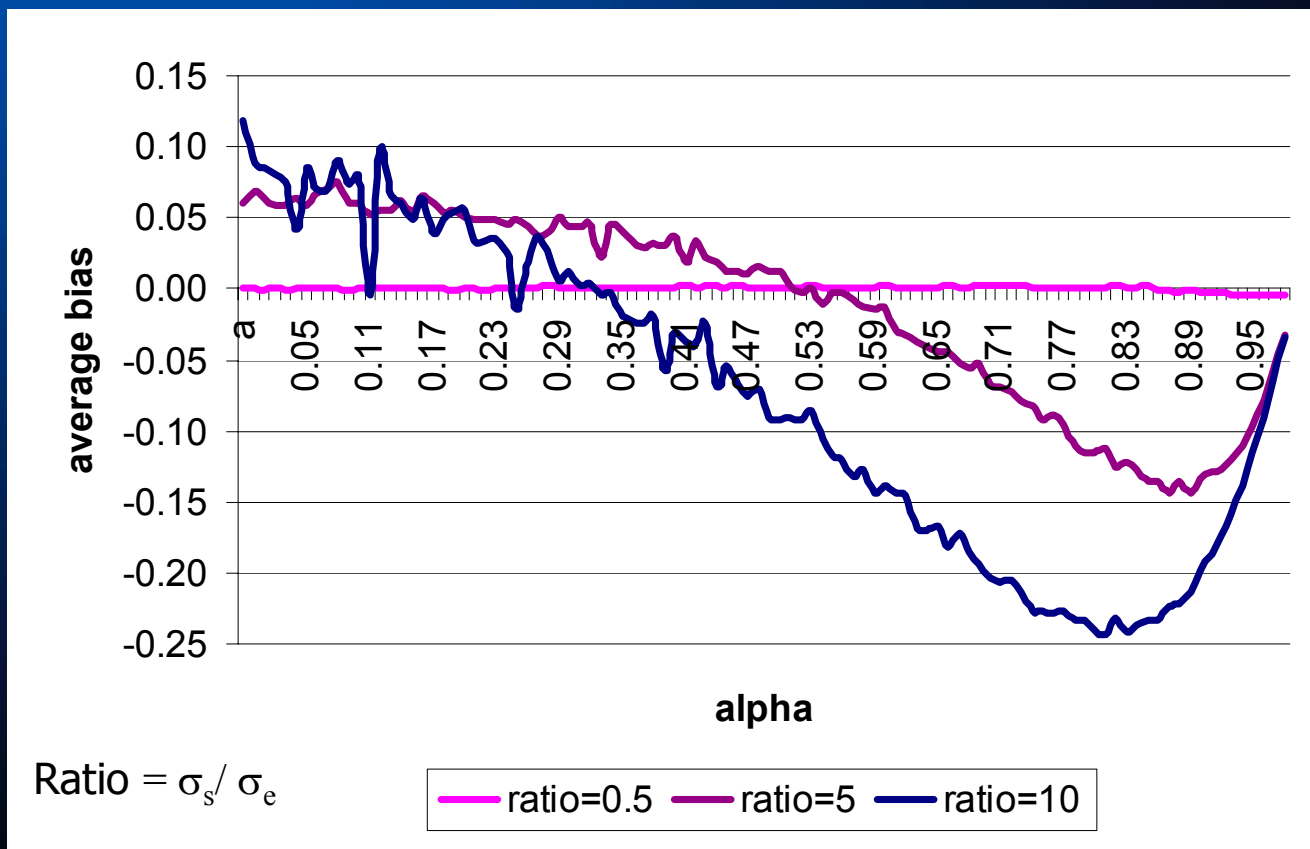
a) Monte-Carlo experiments to evaluate bias in estimates based on the simple model with AR(1) process for the transitory component s_t :

$$\begin{aligned}r_t &= q_t + s_t \\s_t &= \alpha s_{t-1} + e_t \quad 0 \leq \alpha < 1 \quad e_t \sim iidN(0, \sigma_e^2) \\q_t &= q_{t-1} + u_t \quad u_t \sim iidN(0, \sigma_u^2)\end{aligned}$$

In general: tendency to underestimate the adjustment coefficient $(1-\mu)$, and the short run pass-through elasticity

Downward bias increases with the variance of transitory shocks e_t in relation to permanent shocks u_t

- Ambiguous results for the long run pass-through elasticity ($1-\gamma_1$).
- Downward bias peaks at a relatively high degree of serial correlation in s_t

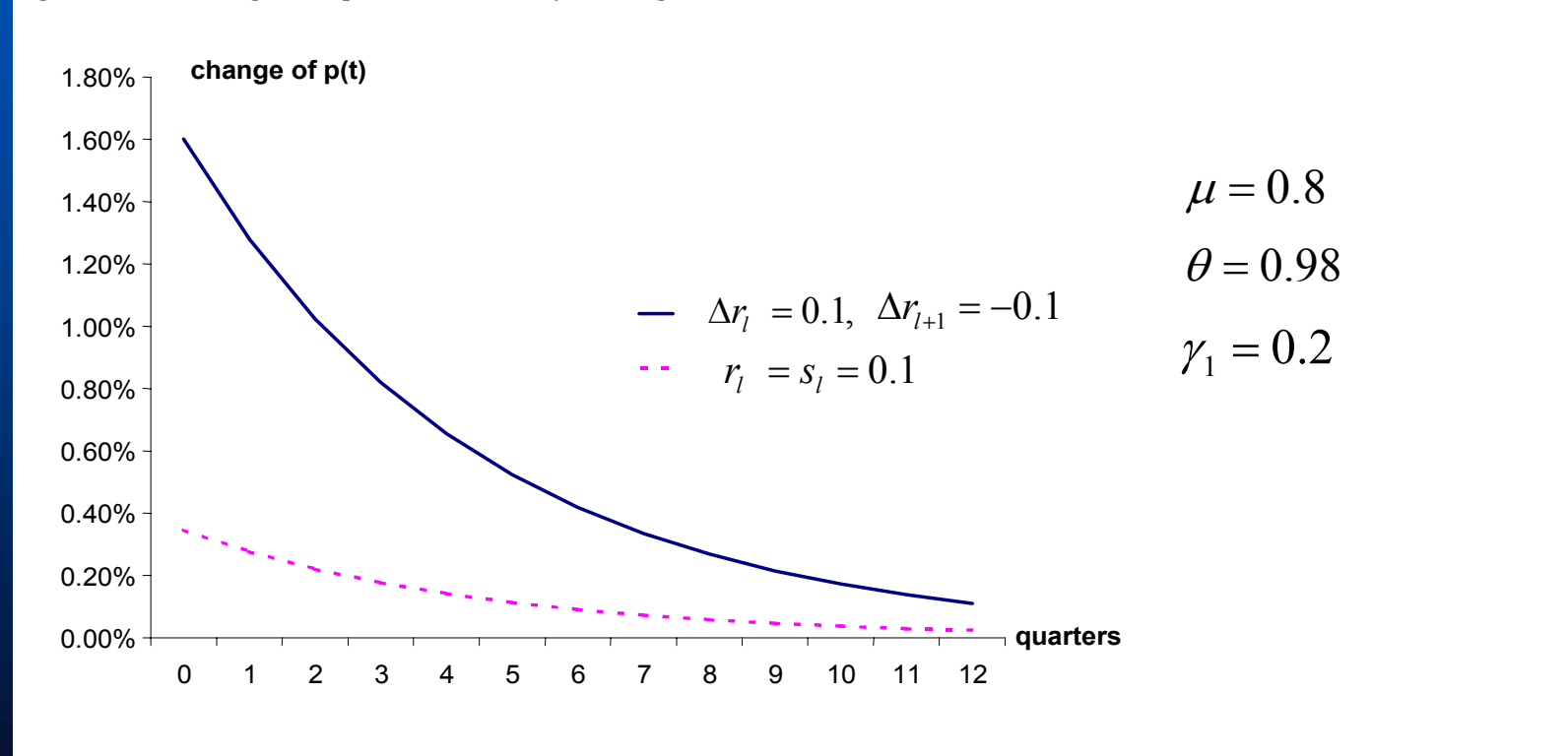


b) simulation of a temporary 10% depreciation at time $t = t$:

1. considering model (10a) and (10b): $r_t = s_t = 0.1$

2. setting $\Delta r_t = 0.1$ and $\Delta r_{t+1} = -0.1$ in conventional ECM

Figure 3: Forecasting the impact of a transitory exchange rate shock



→ Extract transitory component s_t and include contemporaneous and possibly lagged values when estimating the dynamic import supply functions

Bivariate Blanchard-Quah decomposition:

- Joint process of two variables (x_t, z_t) is governed by two shocks
- One variable is stationary \Rightarrow no impact of shocks on long-run level
- Other variable is non-stationary, but one of the shocks is assumed to have no long-run impact

\Rightarrow identification of two independent shocks from VMA representation of $(\Delta x_t, z_t)$

Lee and Chinn (2002): joint process of r_t and current account in relation to nominal GDP (CA_t) satisfy BQ assumptions:

- Two shocks govern the joint process: real demand and supply shocks (ε_t^r) and nominal (financial market) shocks (ε_t^n)
- CA is stationary
- Nominal shocks only have a temporary impact on r_t

Bivariate Blanchard-Quah decomposition

based on the joint process of r_t and current account in relation to nominal GDP (CA_t) - Lee and Chinn (2002):

assumptions:

- Two shocks govern the joint process: real demand and supply shocks (ε_t^r) and nominal (financial market) shocks (ε_t^n)
- CA is stationary
- Nominal shocks only have a temporary impact on r_t

Hypothesized model:
$$\begin{bmatrix} \Delta r \\ CA \end{bmatrix}_t = C(L) \begin{bmatrix} \varepsilon^r \\ \varepsilon^n \end{bmatrix}_t$$

Steps in the decomposition :

1. Unit-root tests for CA_t

→ stationarity cannot be rejected

2. Estimation of the bivariate VAR:
$$\begin{bmatrix} \Delta r \\ CA \end{bmatrix}_t = A(L) \begin{bmatrix} \Delta r \\ CA \end{bmatrix}_{t-1} + \begin{bmatrix} e_1 \\ e_2 \end{bmatrix}_t$$

→ selection of lag order using Akaike and Schwarz Bayesian criterion

3. Transformation into VMA:
$$\begin{bmatrix} \Delta r \\ CA \end{bmatrix}_t = (I - A(L)L)^{-1} \begin{bmatrix} e_1 \\ e_2 \end{bmatrix}_t = (I - A(L)L)^{-1} C(0) \begin{bmatrix} \varepsilon^r \\ \varepsilon^n \end{bmatrix}_t$$

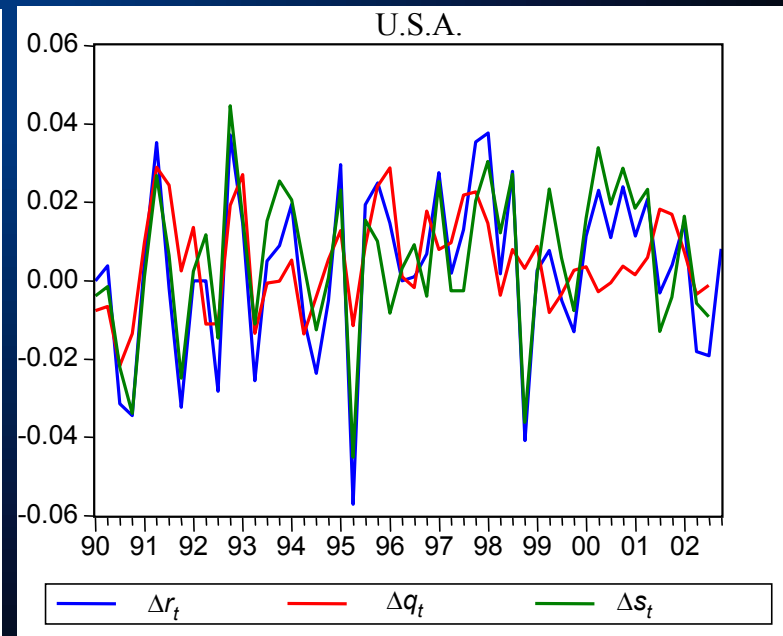
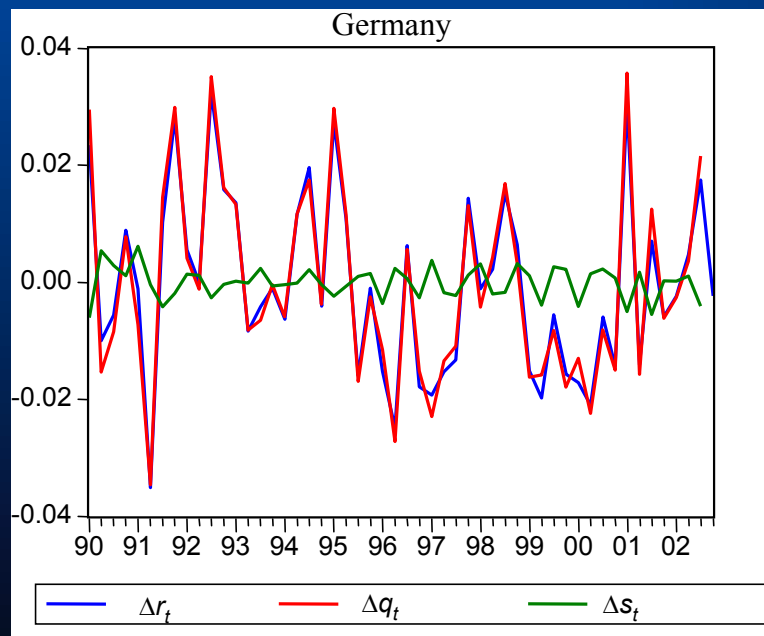
4. Identification of C(0)shocks making use of:
$$[[I - A(1)]C(0)]_{1,2} = 0$$

5. Calculation of transitory and permanent components in r_t

$$\Delta q_t = (1,0) \left[(I - A(L)L)^{-1} C(0) \begin{bmatrix} \varepsilon^r \\ 0 \end{bmatrix}_t \right]$$

$$\Delta s_t = (1,0) \left[(I - A(L)L)^{-1} C(0) \begin{bmatrix} 0 \\ \varepsilon^n \end{bmatrix}_t \right]$$

- For Japan, Germany, France, and Italy real shocks and permanent exchange rate movements dominate
- U.S. real exchange rate is dominated by nominal shocks and transitory changes



Results for the augmented ECMs including transitory exchange rate components s_t, s_{t-1}, \dots :

- Only for the U.S. import price s_{t-1} has a significant impact, for other countries at most s_t
- Transitory component enters with correct (negative) sign
- In many cases increase in short and long run pass-through

	long run ($1-\gamma_1$)	short run
U.S.	0.79 ↑	0.44 ↑
Japan	0.85 ~	0.75 ~
Germany	1.00 ↑	0.63 ~
France	1.38 ↑	0.74 ↑
Italy	1.04 ↑	0.49 ~

- Changes in coefficients are broadly consistent with simulations, exception is adjustment coefficient ($1-\mu$)

- Estimation of ECMs for import supply relationships confirms incomplete pass-through of exchange rates
- Quadratic adjustment cost model highlights role of expected persistence of exchange rate changes
 - implications for estimation and forecasting
- Accounting for permanent and transitory exchange-rate movements in many cases increases ECM estimates of short and long run pass-through
- Predictions of simulations for bias in estimates are not fully supported by the data
 - indicative character of results
 - more extensive study about implications for estimation
 - alternative decompositions